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# Pricing, hedging and calibrating credit from the top down

**Kay Giesecke**

Management Science & Engineering

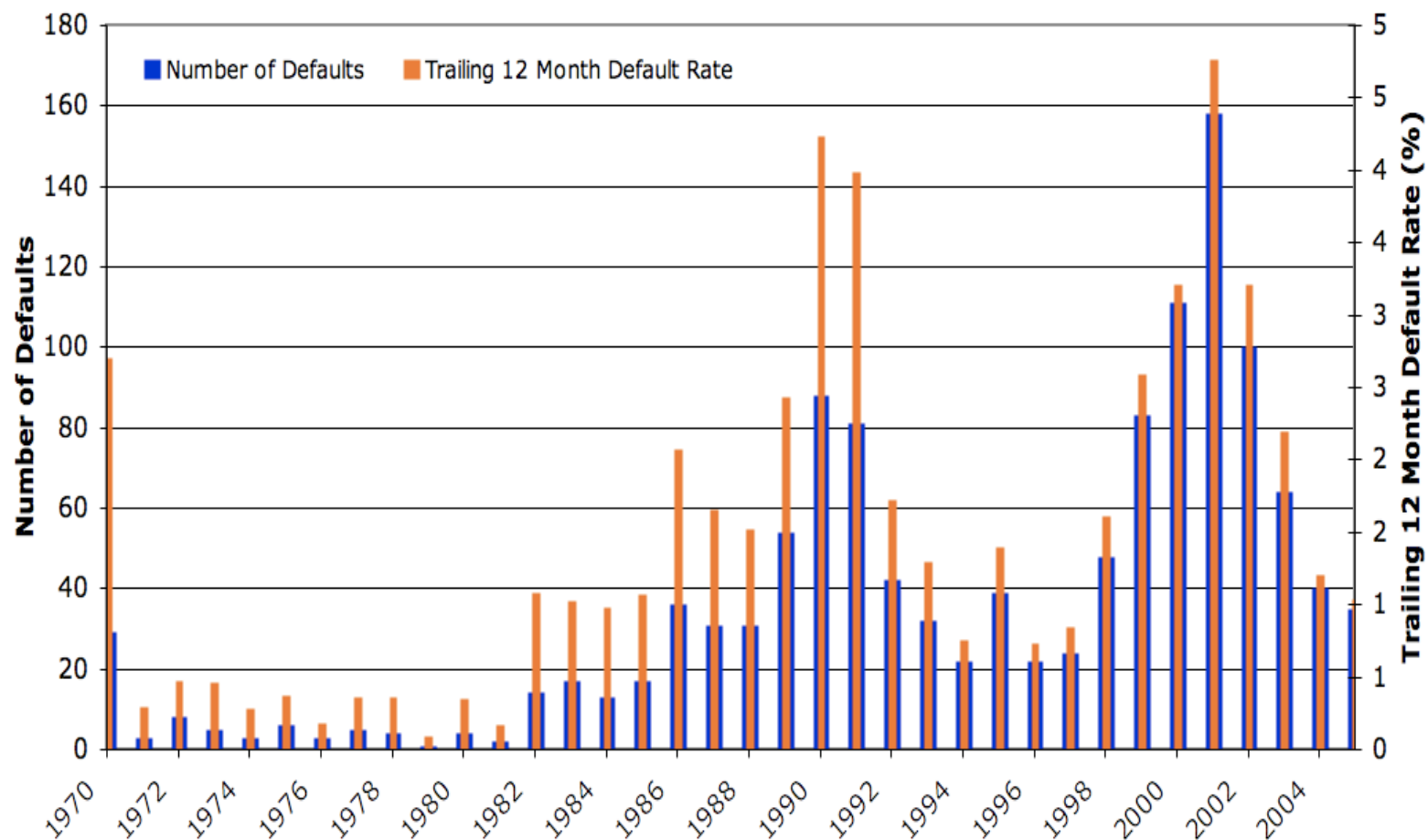
Stanford University

`giesecke@stanford.edu`

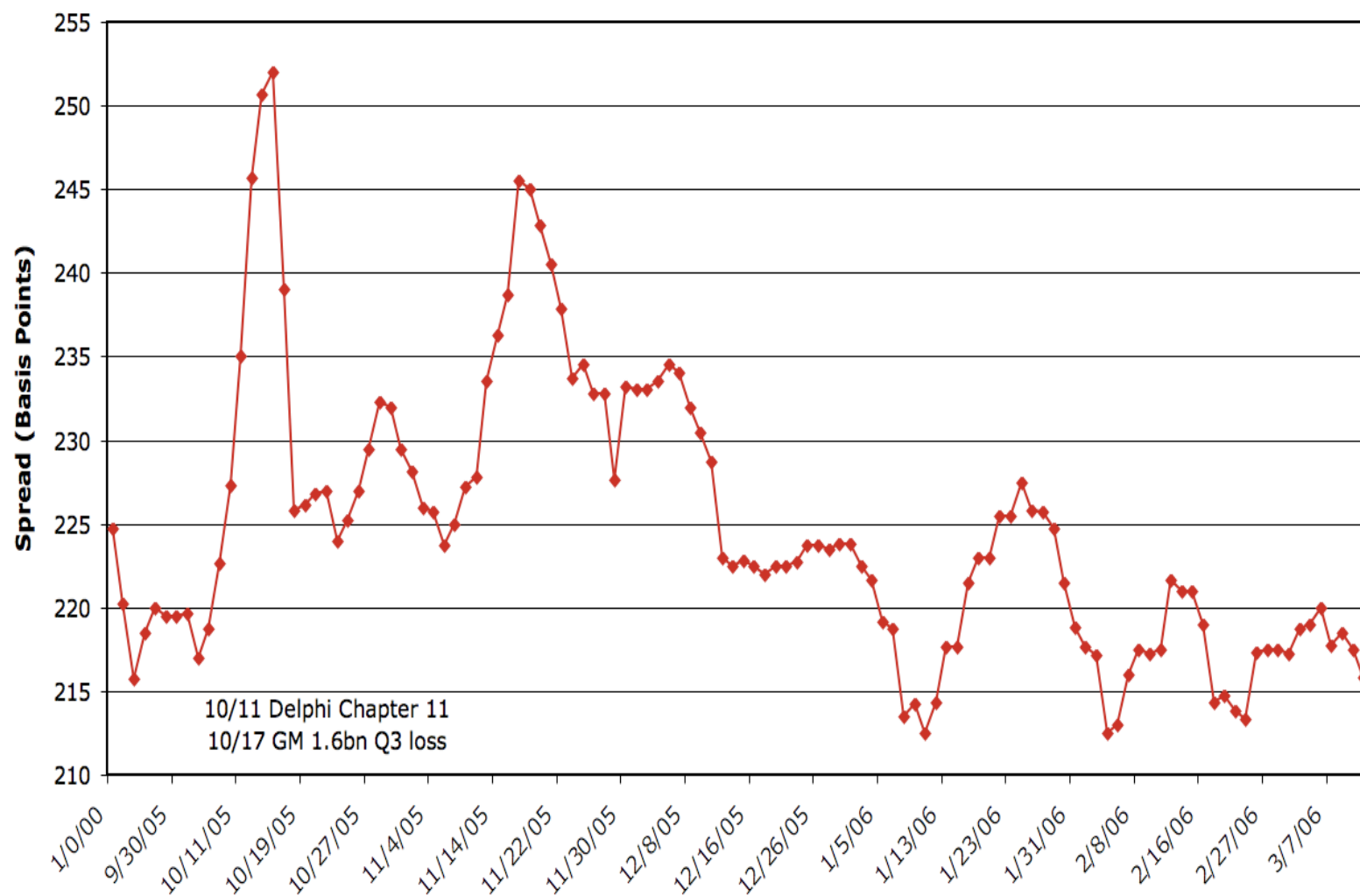
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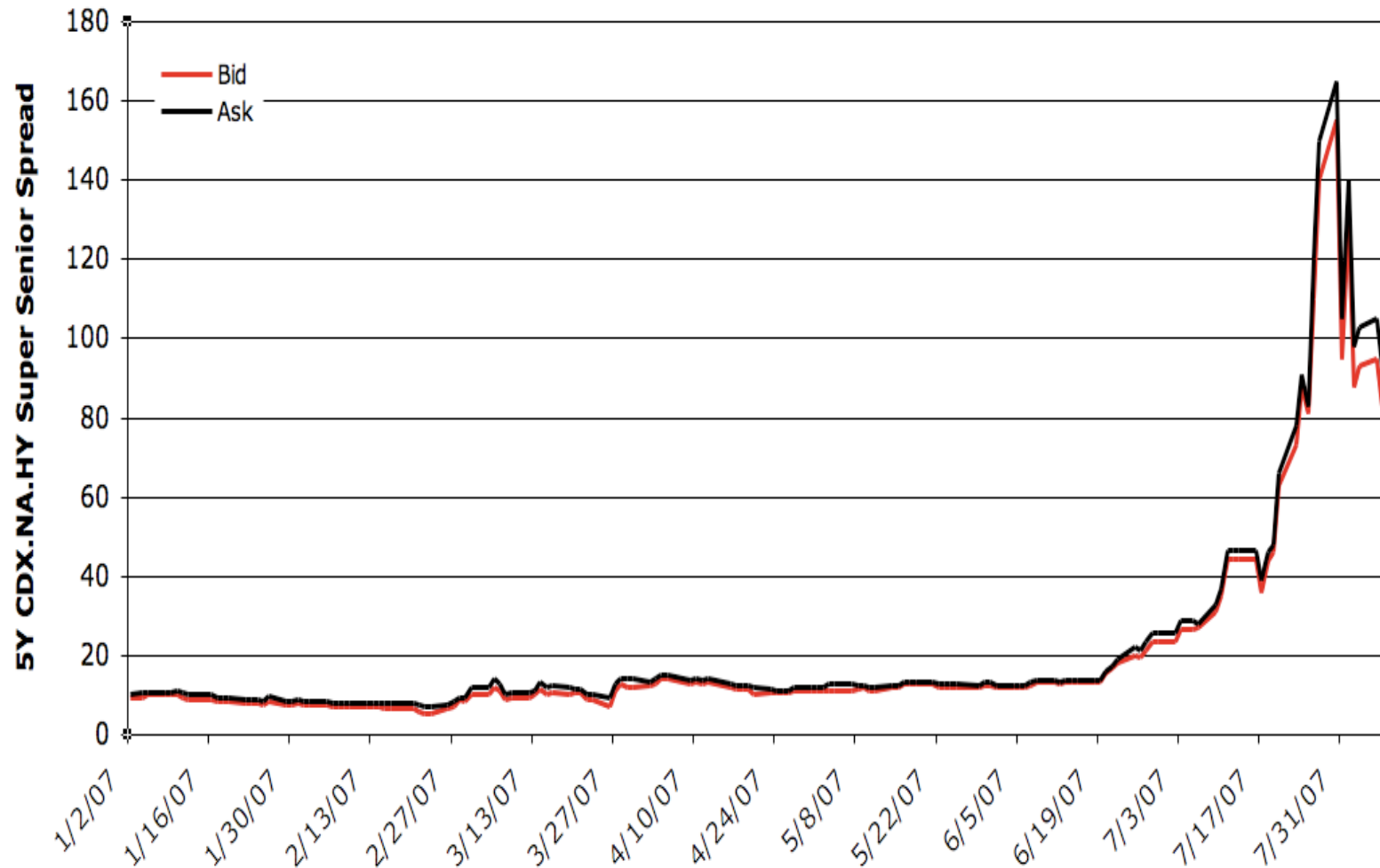
# Defaults cluster: Moody's rated US issuers



# Feedback in 5Y CDX.NA.XO Index



# Feedback in 5Y CDX.NA.HY Super senior



# Default dependence

- **Cyclical correlation:** first order effect
  - Firm sensitivity to common economic factors
- **Feedback or contagion:** second order effect
  - Complex web of business, informational and legal relationships
  - Important for contracts on loss volatility (e.g. tranche option)
  - Important for hedging of single name spread and default risk
  - Important for super senior risk

# Portfolio credit derivative

- Trading of default dependence in a portfolio of names
  - Targets portfolio loss or volatility of loss
- Contingent claim on portfolio loss
  - Defaults occur at stopping times  $T^1 < T^2 < \dots$
  - **Default process**  $N$  counts defaults

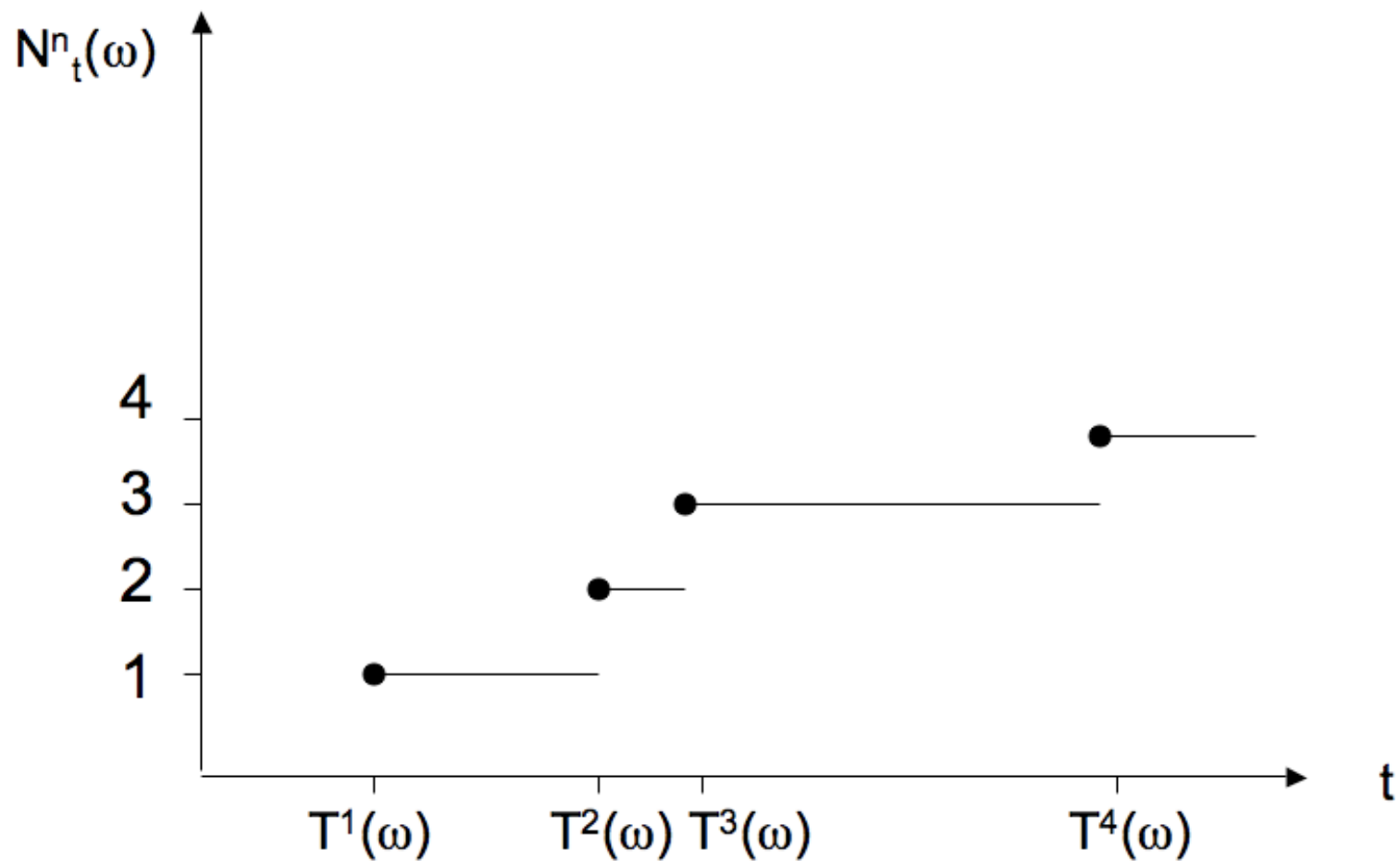
$$N_t = \sum_{k: T^k \leq t} 1$$

- **Loss process**  $L$  records financial loss due to default

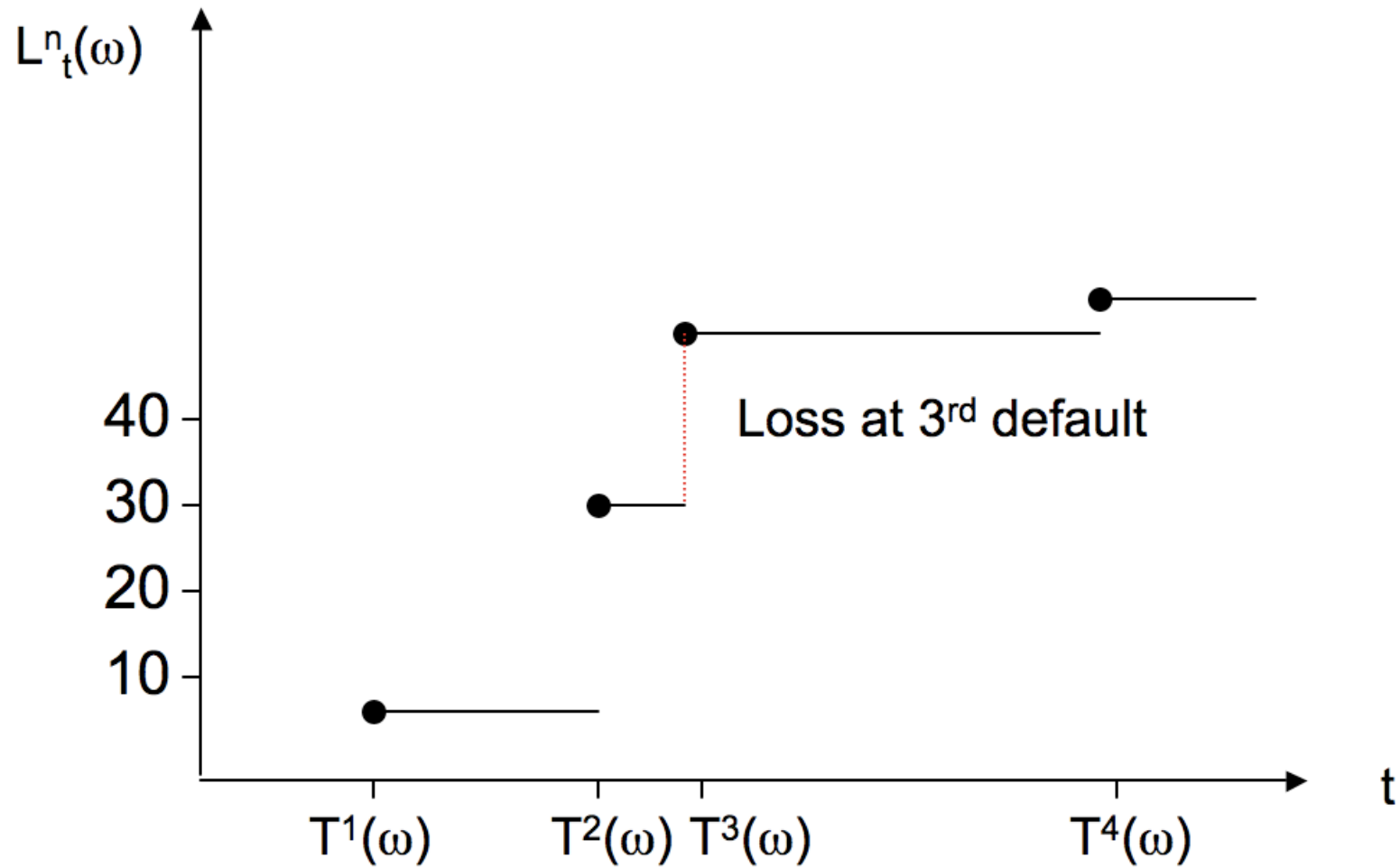
$$L_t = \sum_{k: T^k \leq t} \ell^k$$

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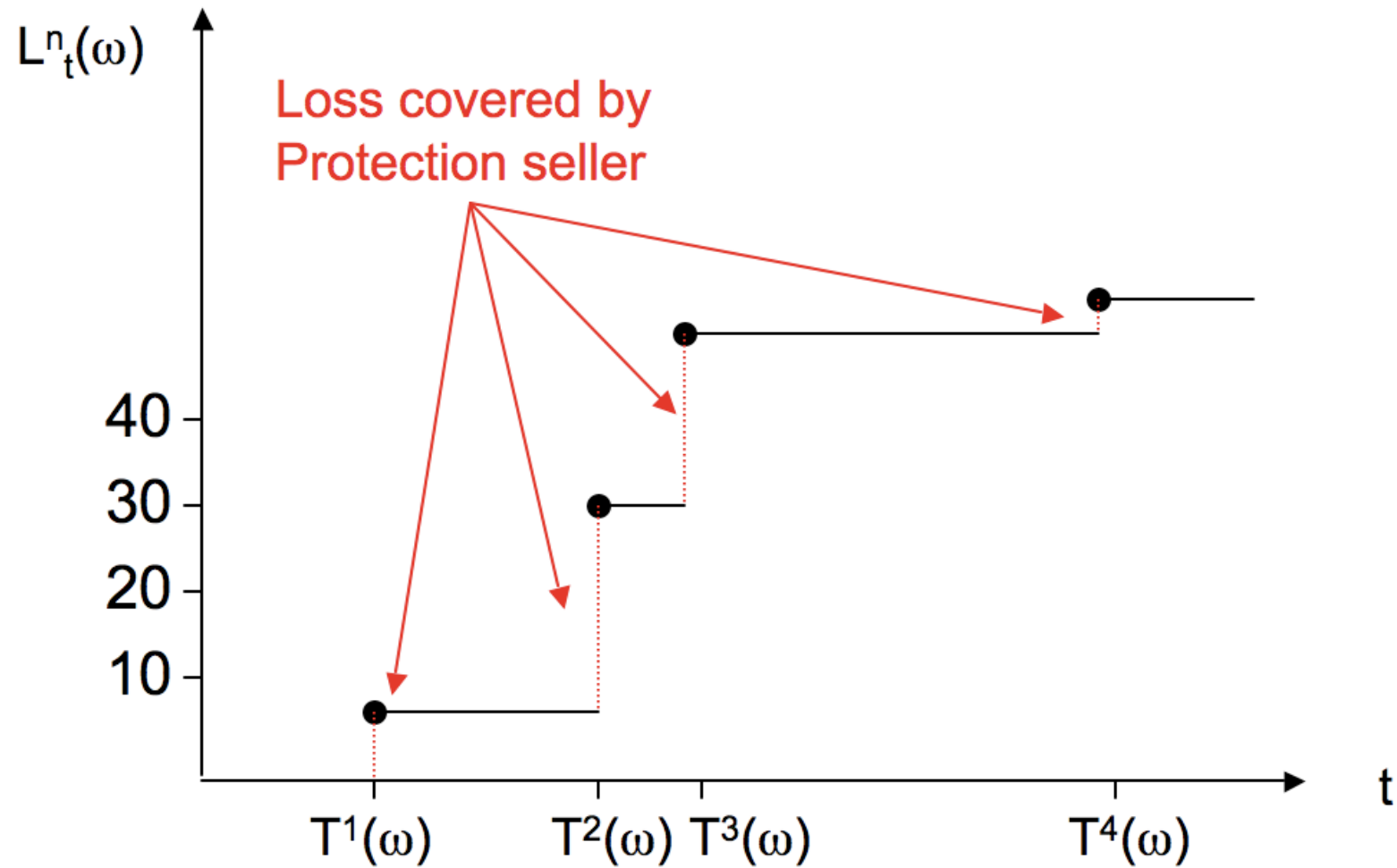
# Sample path of default process



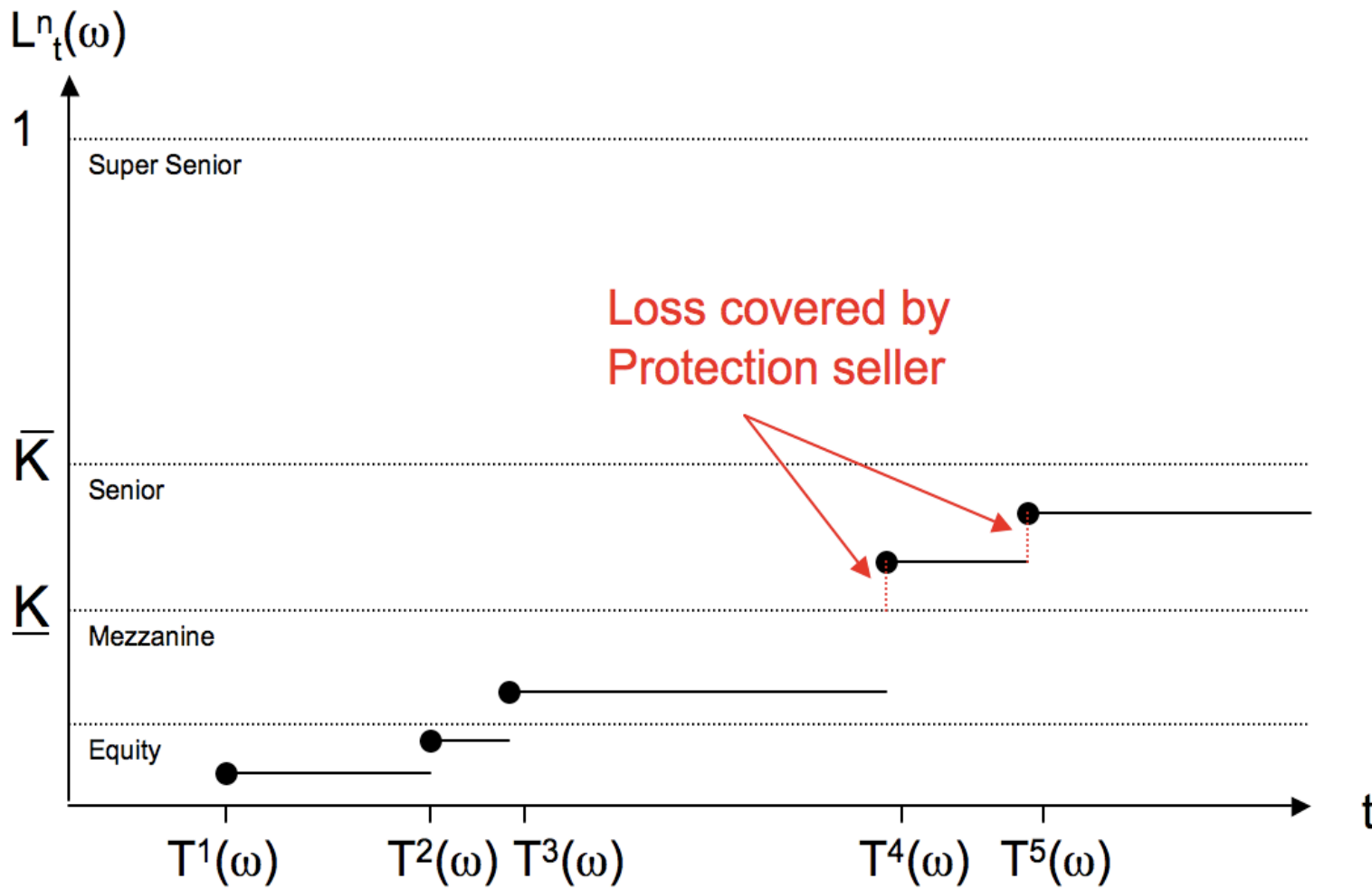
# Sample path of loss process



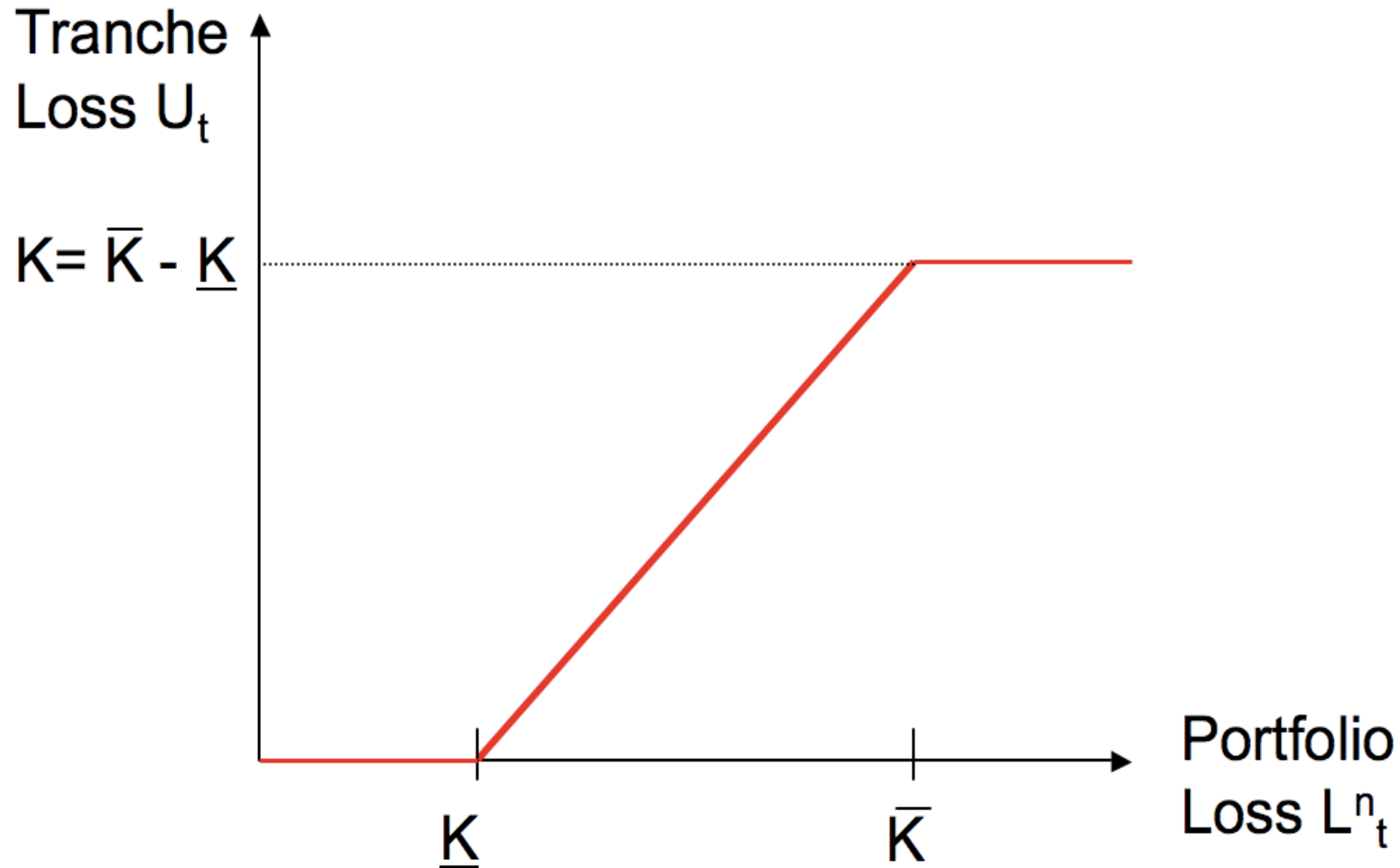
# Loss process and index swap



# Loss process and tranche swap



# Tranche loss = call spread on portfolio loss



## Tranche swap

- **Default leg** covers the tranche loss  $U_t = (L_t - \underline{K})^+ - (L_t - \overline{K})^+$  up to the maturity  $T$ ; with a constant interest rate  $r$

$$D_t = e^{-r(T-t)} E_t[U_T] - U_t + r \int_t^T e^{-r(s-t)} E_t[U_s] ds$$

- **Premium leg**: stream of payments at  $(t_m)$  that are proportional to  $K - U_{t_m}$ ; with  $S$  the spread

$$P_t(S) = S \sum_{t_m \geq t} e^{-r(t_m-t)} (K - E_t[U_{t_m}])$$

- A top down value of the **tranche spread** at time  $t$  is the solution  $S = S_t(T)$  to the equation  $D_t = P_t(S)$ 
  - Depends only on the values of call options on  $L$

# Top down approach

- The loss process  $L$  is specified directly in terms of
  - An **intensity**  $\lambda$  that governs the arrival of events: for small  $\Delta t$

$$\lambda_t \Delta t \approx P_t[\text{event in } (t, t + \Delta t)]$$

- A **distribution**  $\nu$  that governs the random loss at default  $\ell^k$
- A portfolio derivative is a contingent claim on  $L$ 
  - Basic building block is the call option  $E_t[(L_s - k)^+]$
- Random thinning allocates  $L$  to the portfolio constituents
  - Hedging of single name exposures

# Compound Poisson loss process

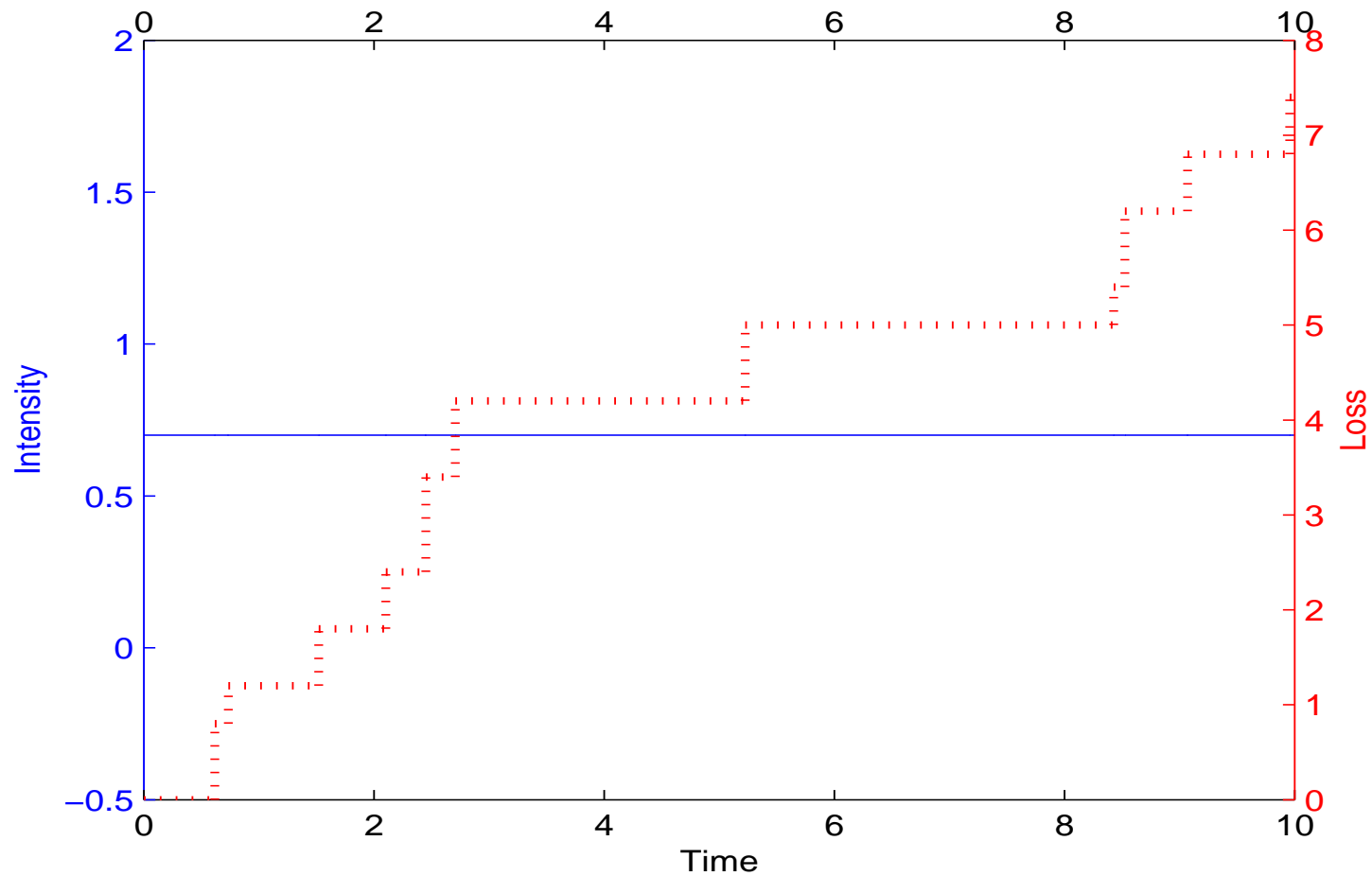
- The intensity  $\lambda$  is given by

$$\lambda_t = c = \text{constant}$$

- Waiting times between events are independent exponential variables with parameter  $c$
  - The benchmark top down model for the portfolio loss
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# Compound Poisson loss process

The loss  $\ell^k$  is uniform on  $\{0.4, 0.6, 0.8, 1\}$  and  $c = 0.7$



# Compound birth loss process

- The intensity  $\lambda$  is given by

$$\lambda_t = c + \delta L_t$$

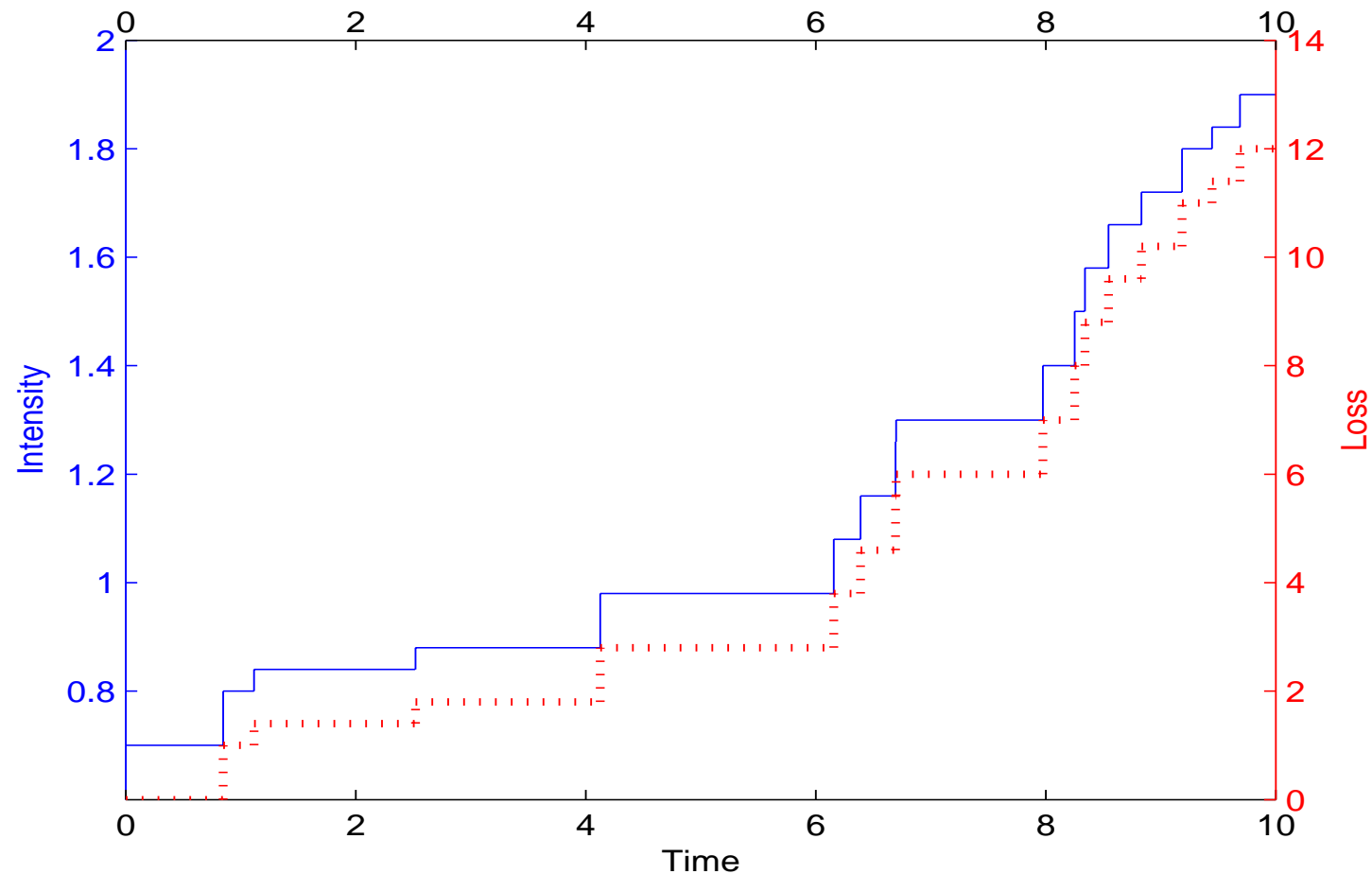
- Intensity responds to events; if  $\delta = 0$ : Poisson process
- **Negative correlation** between default and recovery rates
- Waiting times between events are independent exponentials with (state-dependent) parameter  $c, c + \delta \ell^1, c + \delta(\ell^1 + \ell^2), \dots$
- The intensity has dynamics

$$d\lambda_t = \delta dL_t, \quad \lambda_0 = c$$

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# Compound birth loss process

The loss  $\ell^k$  is uniform on  $\{0.4, 0.6, 0.8, 1\}$ ,  $c = 0.7$  and  $\delta = 0.1$



# Compound Hawkes loss process

- The intensity  $\lambda$  is given by

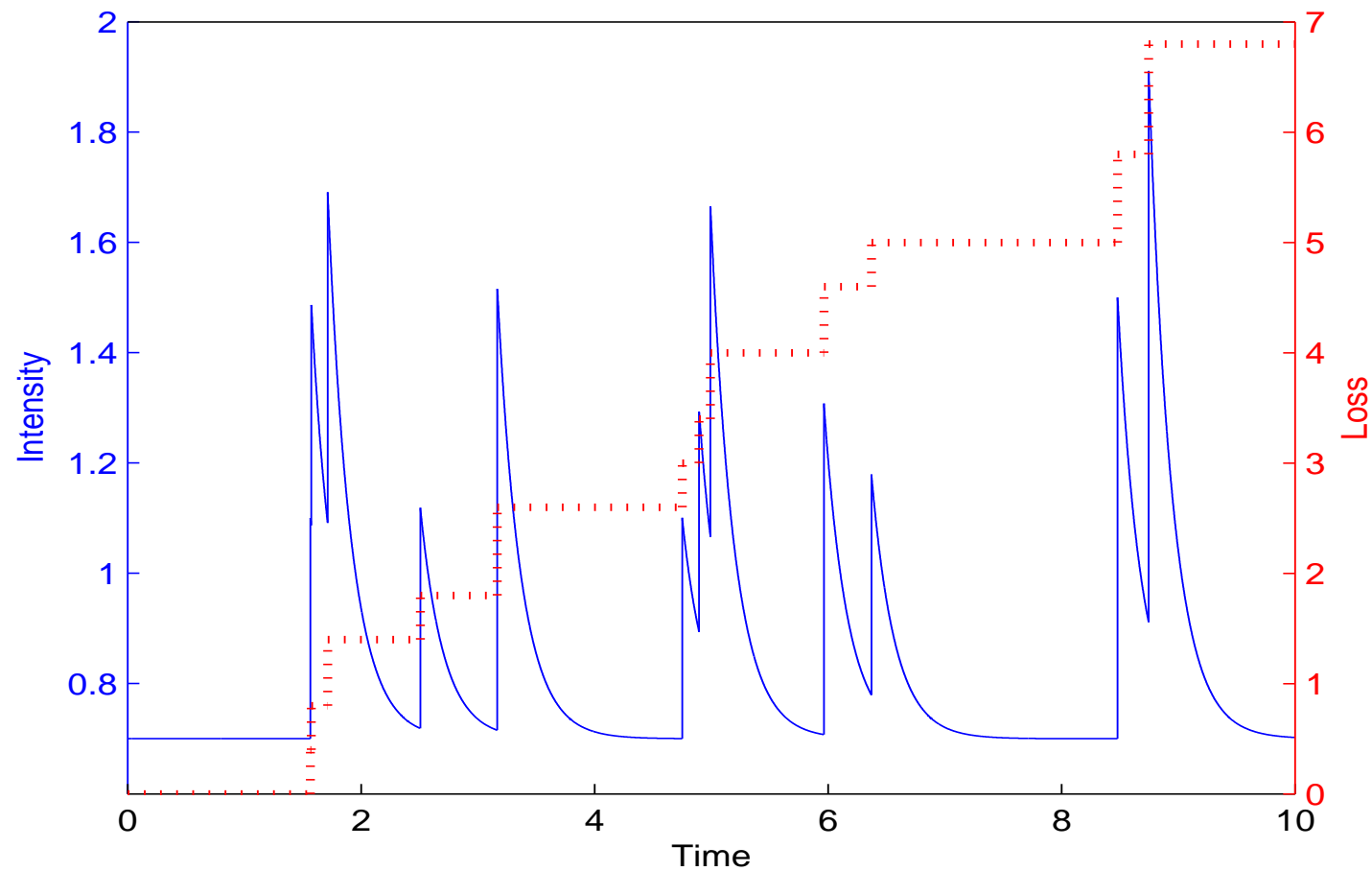
$$\lambda_t = c + (\lambda_0 - c)e^{-\kappa t} + \delta \int_0^t e^{-\kappa(t-s)} dL_s$$

- Intensity responds to events; if  $\delta = \kappa = 0$ : Poisson process
- Intensity decays after an event; if  $\kappa = 0$ : birth process
- Negative correlation between default and recovery rates
- The intensity has dynamics

$$d\lambda_t = \kappa(c - \lambda_t) dt + \delta dL_t, \quad \lambda_0 > 0$$

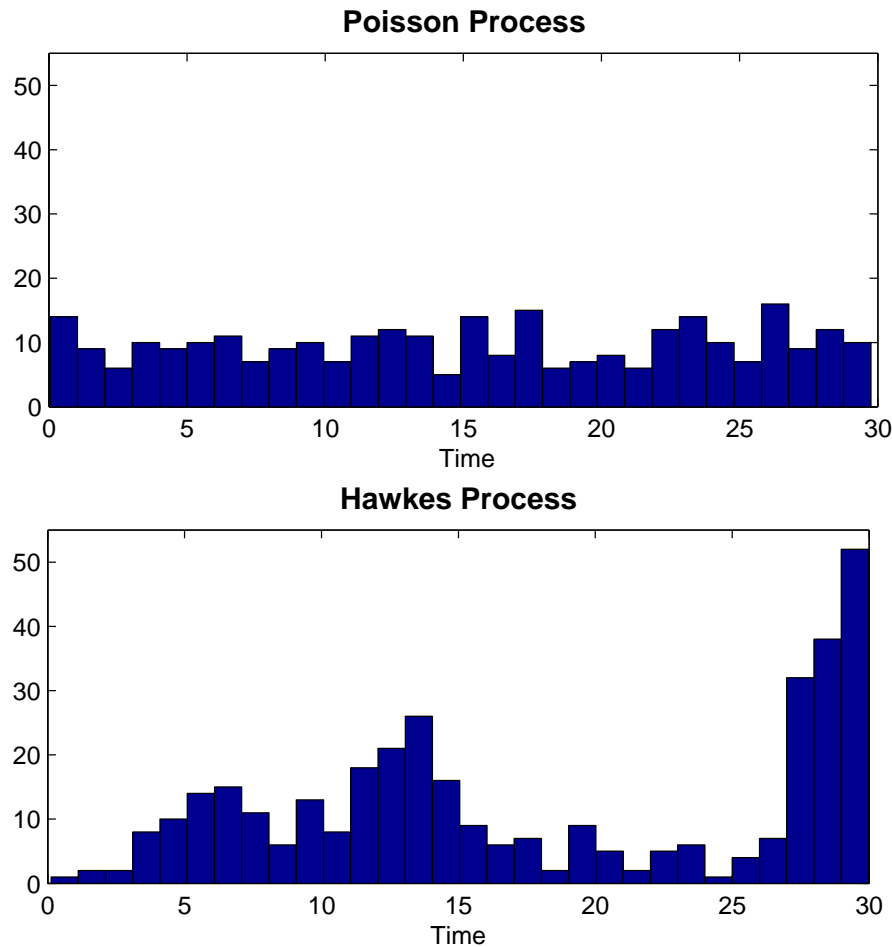
# Compound Hawkes loss process

The loss  $\ell^k$  is uniform on  $\{0.4, 0.6, 0.8, 1\}$ ,  $c = \lambda_0 = 0.7$ ,  $\delta = 1$ ,  $\kappa = 5$



# Hawkes vs. Poisson models

The loss  $\ell^k$  is uniform on  $\{0.4, 0.6, 0.8, 1\}$ ,  $c = \lambda_0 = 1$ ,  $\delta = 2$  and  $\kappa = 1.5$  for the Hawkes process and  $c = 10.57$  for the Poisson process



## Basic affine loss process

- The intensity  $\lambda$  is given by

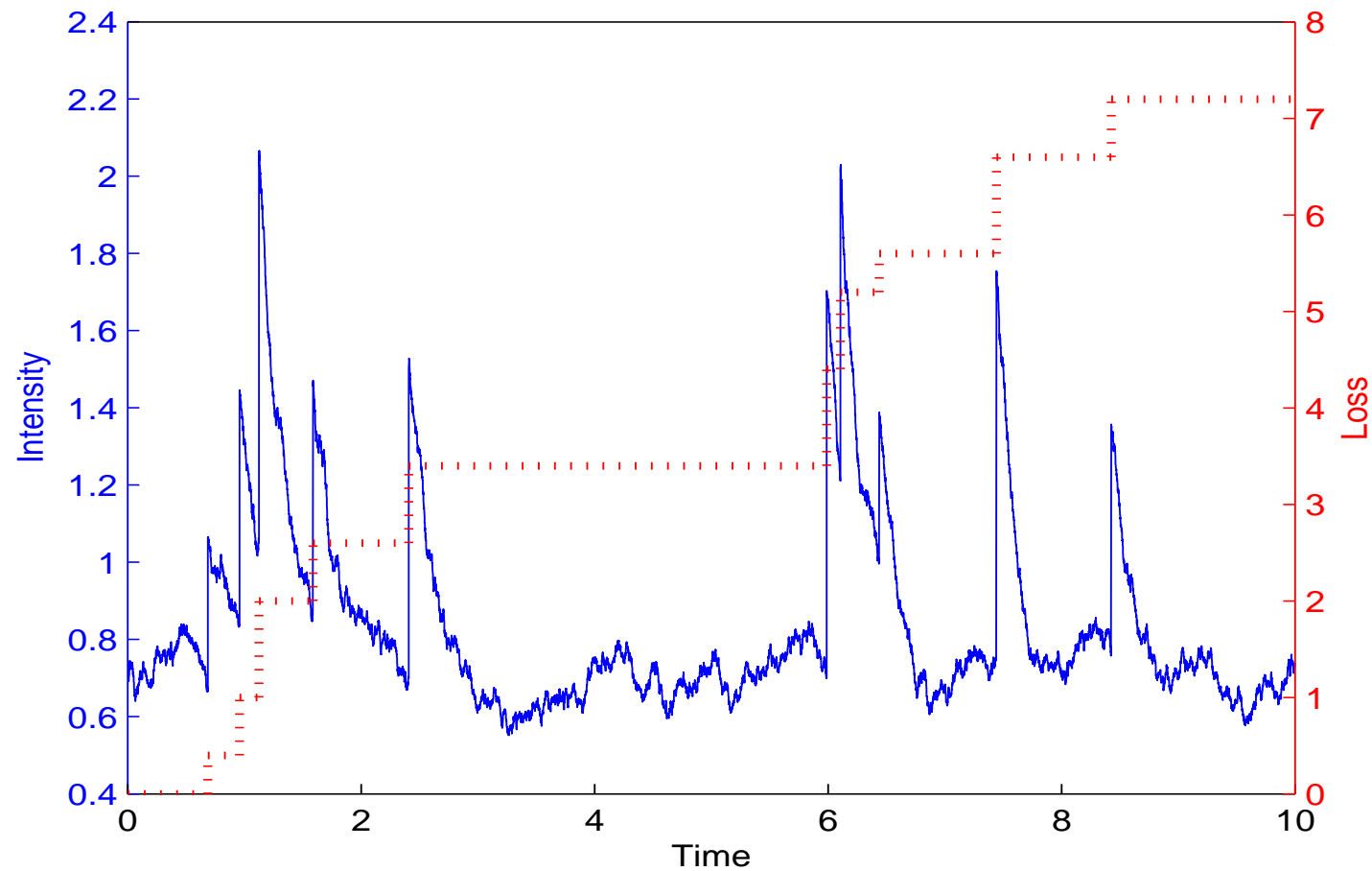
$$\lambda_t = c + (\lambda_0 - c)e^{-\kappa t} + \sigma \int_0^t e^{-\kappa(t-s)} \sqrt{\lambda_s} dW_s + \delta \int_0^t e^{-\kappa(t-s)} dL_s$$

- Intensity diffuses between events; if  $\sigma = 0$ : Hawkes process
- Intensity responds to events; if  $\delta = \sigma = \kappa = 0$ : Poisson process
- Intensity decays after an event; if  $\kappa = \sigma = 0$ : birth process
- Negative correlation between default and recovery rates
- The intensity has dynamics

$$d\lambda_t = \kappa(c - \lambda_t) dt + \sigma \sqrt{\lambda_t} dW_t + \delta dL_t, \quad \lambda_0 > 0$$

# Basic affine loss process

The loss  $\ell^k$  is uniform on  $\{0.4, 0.6, 0.8, 1\}$ ,  $c = \lambda_0 = 0.7$ ,  $\delta = 1$ ,  $\kappa = 5$



# Calibration

- We calibrate the basic affine loss process to index and tranche spreads/upfront rates on the CDX.NA.HY.5Y observed on 5/11/2007, which has attachment points 0, 10, 15, 25, 35, 100%
  - We assume the distribution of loss at default  $\nu$  is uniform on  $\{\ell_1, \ell_2\}$  with  $0 < \ell_1 < \ell_2 < 1$ , and we set the expected loss at default  $E[\ell^k] = \int z d\nu(z) = 0.6$
  - The intensity  $\lambda$  is specified by the parameters  $\lambda_0, c, \kappa, \sigma$  and  $\delta$
  - The risk-free rate  $r = 0.05$
-

# Calibration

- We fit the parameter vector  $\theta = (\lambda_0, c, \kappa, \sigma, \delta, \ell_1)$  by numerically solving the optimization problem

$$\min_{\theta \in \Theta} \sum_i \left( \frac{\text{MarketMid}(i) - \text{Model}(i, \theta)}{\text{MarketAsk}(i) - \text{MarketBid}(i)} \right)^2$$

subject to  $2\kappa c \geq \sigma^2$

where  $\Theta = [0, 5]^3 \times [0, 1] \times [0, 5] \times [0.2, 0.6]$  and the sum ranges over the spread/upfront rate data points

- We use adaptive simulated annealing
- We analyze two model specifications:
  - Mod 1 is unrestricted
  - Mod 2 is restricted: the diffusive volatility  $\sigma = 0$

# Calibration

5Y maturity: market data and fitting results

- Both models fit the data well; the basic affine model does better than the Hawkes model

	MarketBid	MarketAsk	Mod 1	Mod 2
0-10	70.50%	70.75%	71.11%	71.48%
10-15	34.25%	34.50%	32.85%	32.74%
15-25	316.00	319.00	316.80	311.43
25-35	79.00	81.00	81.47	77.34
Index	262.85	263.10	263.46	262.97
MinObj			41.63	60.41
AAPE			1.47%	2.24%

# Calibration

5Y maturity: initial and calibrated parameter values

- We ran several calibrations with different initial values; the values reported below generated the lowest objective function value
- The calibrated parameter values are very similar for the two models

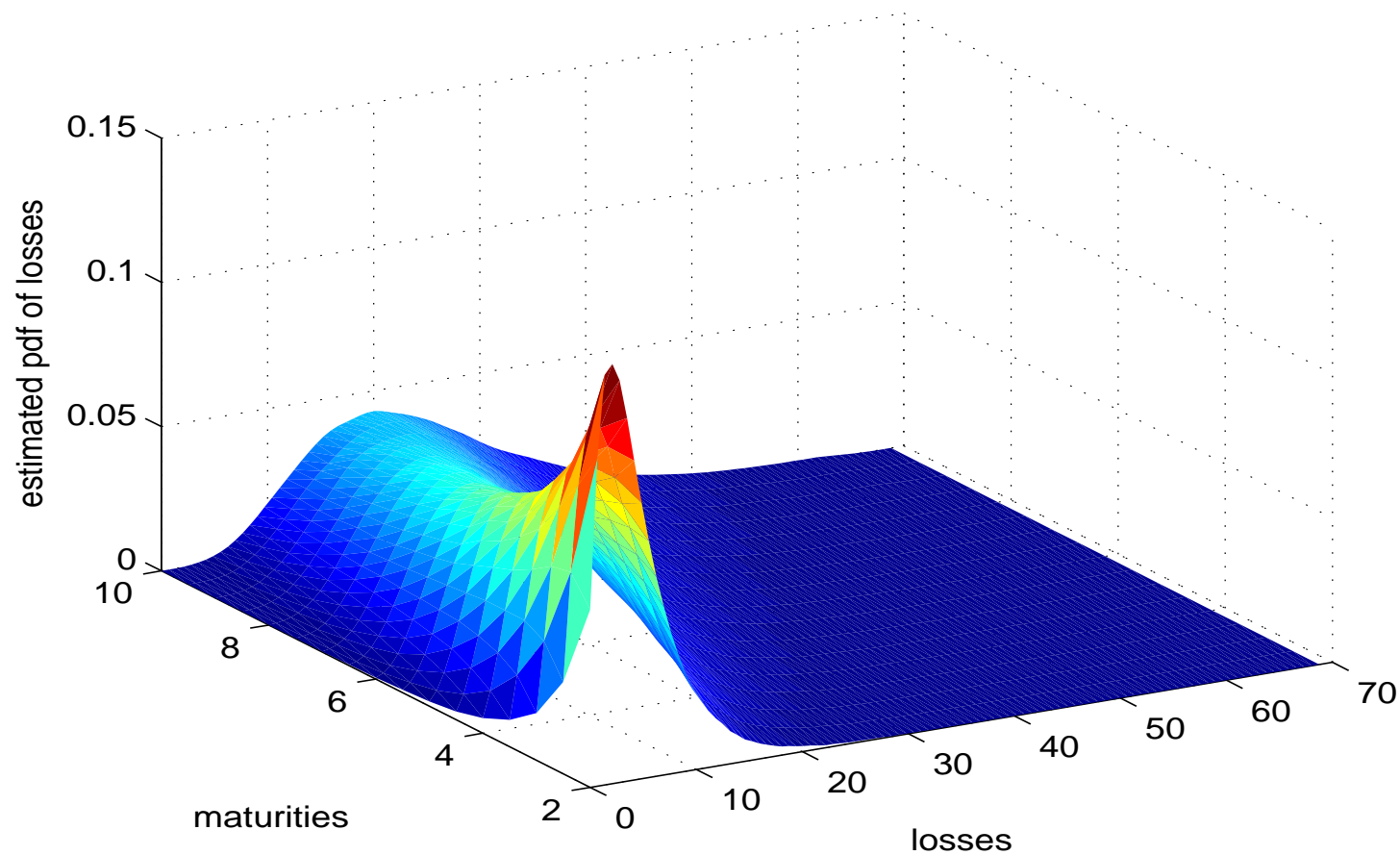
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	$\lambda_0$	$c$	$\kappa$	$\sigma$	$\delta$	$\ell_1$
Initial	2.50	2.50	2.50	0.50	2.50	0.40
Mod 1	0.70	1.61	2.62	0.62	2.99	0.24
Mod 2	0.75	1.60	2.58	0.00	2.94	0.24

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# Calibration

5Y maturity: loss distribution implied by the calibrated basic affine model Mod 1



# Calibration

5Y maturity: loss distribution implied by the calibrated model

- The loss distribution for all horizons is captured by one set of parameters  $\theta = (\lambda_0, c, \kappa, \sigma, \delta, \ell_1)$ , and so is the complete term structure of arbitrage free index and tranche spreads for all attachment points
    - Compare with copula models, which need one correlation parameter per maturity or attachment point
  - Further, we have the evolution of the conditional loss distribution for all horizons as the observation date moves forward
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# Calibration

5Y maturity: parameter stability

- We re-calibrate at different dates starting with 5/11/07
- The initial values at a date after 5/11 are set to the optimal values from the previous date

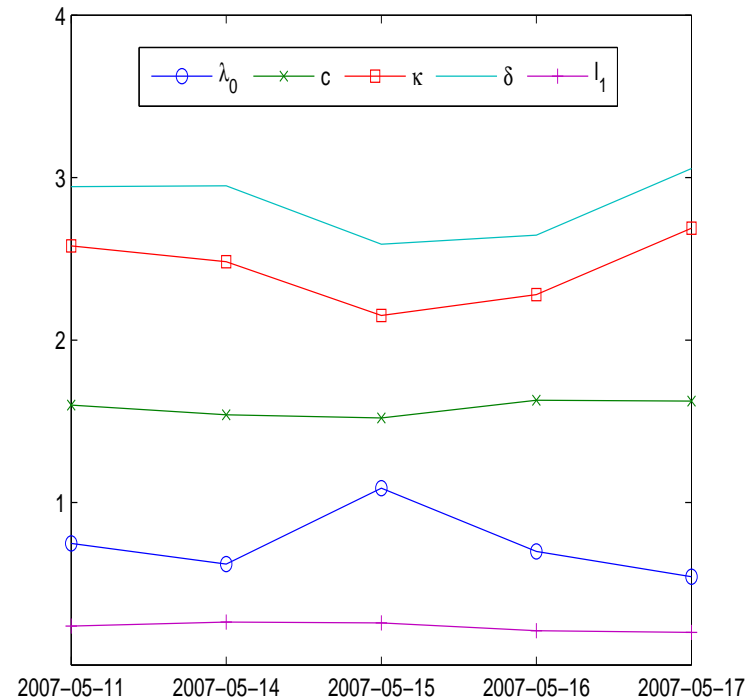
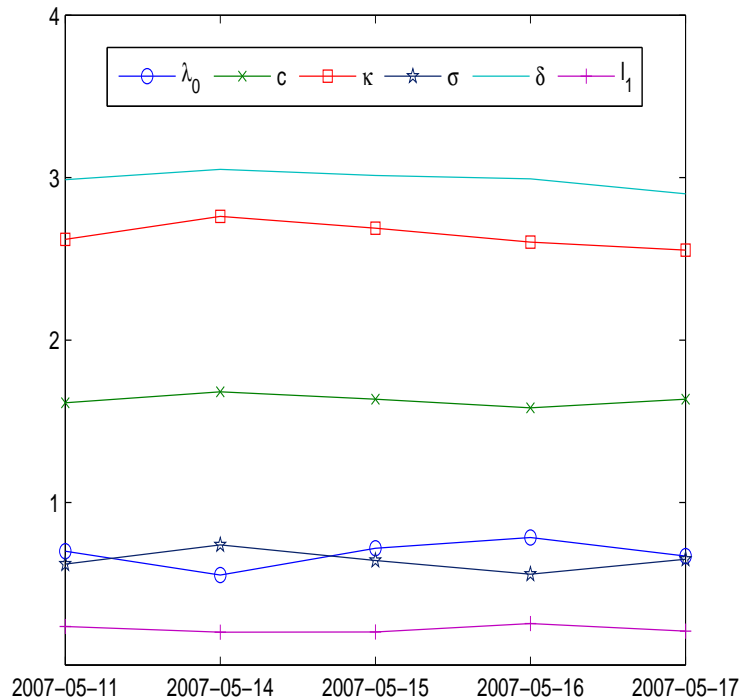
Mod 1	Date	05/11	05/14	05/15	05/16	05/17
	MinObj	41.63	58.66	46.78	46.24	58.24
	AAPE	1.47%	1.70%	1.26%	1.32%	1.35%

Mod 2	Date	05/11	05/14	05/15	05/16	05/17
	MinObj	60.41	73.43	54.29	43.88	65.23
	AAPE	2.24%	2.04%	1.57%	1.45%	1.92%

# Calibration

5Y maturity: parameter stability Mod 1 (left) better than Mod 2



# Calibration

5 and 7Y maturities: market data and fitting results

		MarketBid	MarketAsk	Mod 1	Mod 2
5Y	0-10	70.50%	70.75%	71.69%	72.00%
	10-15	34.25%	34.50%	33.44%	33.47%
	15-25	316.00	319.00	316.13	309.93
	25-35	79.00	81.00	78.45	72.79
	Index	262.85	263.10	263.94	262.75
7Y	0-10	80.13%	80.38%	81.23%	81.49%
	10-15	55.50%	55.75%	55.17%	55.40%
	15-25	582.00	587.00	580.25	584.63
	25-35	180.00	183.00	206.41	207.98
	Index	307.50	307.75	307.53	308.91
	MinObj			136.12	192.82
	AAPE			2.35%	3.30%

# Calibration

5Y maturity: Mod 1 out-of-sample forecast for 7Y

		MarketBid	MarketAsk	5+7Y	5Y
5Y	0-10	70.50%	70.75%	71.69%	71.11%
	10-15	34.25%	34.50%	33.44%	32.85%
	15-25	316.00	319.00	316.13	316.80
	25-35	79.00	81.00	78.45	81.47
	Index	262.85	263.10	263.94	263.46
7Y	0-10	80.13%	80.38%	81.23%	81.79%
	10-15	55.50%	55.75%	55.17%	52.49%
	15-25	582.00	587.00	580.25	531.61
	25-35	180.00	183.00	206.41	186.26
	Index	307.50	307.75	307.53	296.00
	MinObj			136.12	
	AAPE			2.35%	3.04%

# Single name hedging

## Random thinning

- Models for the default times  $\tau_i$  of the portfolio constituents are required to estimate the exposure to single name spread changes
  - We need to attribute a fraction of the portfolio credit risk to a constituent name
  - The portfolio arrival risk is described by the intensity  $\lambda$  so we allocate a fraction  $Z^i$  of  $\lambda$  to name  $i$
  - The fraction  $Z^i$  is parametrized and fit so that we match the single name swap term structure
  - The fraction  $Z^i$  may follow its own stochastic process, and is called the **thinning** for  $\lambda$
-

# Single name hedging

## Single name default probabilities

- For any model  $\lambda$ , there exists a thinning process  $Z^i$  such that

$$P_t[t < \tau_i \leq s] = \int_t^s E_t[Z_u^i \lambda_u] du$$

- The thinning process satisfies

$$Z_t^i = \lim_{\epsilon \rightarrow 0} Z_t^i(\epsilon)$$

where  $Z_t^i(\epsilon)$  is the conditional probability at time  $t$  that the next defaulter is  $i$ , given that a default occurs by time  $t + \epsilon$ :

$$Z_t^i(\epsilon) = \sum_k \frac{P_t[T^k = \tau_i \text{ and } T^k \leq t + \epsilon]}{P_t[T^k \leq t + \epsilon]} \mathbf{1}_{\{T^{k-1} \leq t < T^k\}}$$

# Single name hedging

Properties of a thinning process

1. Each  $Z_t^i$  is a conditional probability, so  $Z_t^i \in [0, 1]$  for all  $i$  and  $t$
  2. Unless the whole portfolio is in default, the  $Z_t^i$  must sum to 1 over constituents  $i$
  3. Whenever the whole portfolio is in default the  $Z_t^i$  must vanish
-

# Single name hedging

Example specification

- Let  $n$  be the number of names in the portfolio
- Suppose the thinning processes  $Z^i$  are given by

$$Z_t^i = e^{z^i(t)} 1_{\{N_t \leq n\}}$$

for deterministic functions that satisfy

$$z^i(t) \leq 0, \quad \sum_{i=1}^n e^{z^i(t)} = 1$$

- Therefore

$$P_t[t < \tau_i \leq s] = \int_t^s e^{z^i(u)} E_t[\lambda_u 1_{\{N_u \leq n\}}] du$$

which can be calculated for all the models  $\lambda$  discussed above

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# Single name hedging

## Single name CDS

- Given the default probability  $F^i(t) = P_0[\tau_i \leq t]$ , we can price a CDS referenced on firm  $i$ , or other single name derivatives
- If the loss at default is paid at the default, the premium payments occur at times  $(t_m)$  and we neglect premium accruals, then the spread on a credit swap with maturity  $t$  is given by the formula

$$S^i(t) = \frac{\ell(e^{-rt}F^i(t) + r \int_0^t e^{-rs}F^i(s)ds)}{\sum_{t_m} e^{-rt_m}(1 - F^i(t_m))}$$

where  $\ell$  is the expected loss at default  $E[\ell^k]$

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# Single name hedging

## Example specification

- If swaps with maturities  $0 = T_0 < T_1 < T_2 < \dots < T_J$  are available, we can choose the  $z^i(u)$  constant between maturities:

$$z^i(u) = z^{ij}, \quad T_{j-1} \leq u < T_j, \quad j = 1, \dots, J$$

- The default probability  $F^i(t)$  can then be calculated as a finite sum so the swap spread is easily computed as well

# Single name hedging

## Calibrating the thinning

- Two-step procedure
    - Fit the parameters of the intensity  $\lambda$  using index and tranche spreads as described above
    - Given the optimal parameter values for  $\lambda$ , separately fit  $\{z^i(u)\}_i$  to single name spreads using a standard bootstrapping algorithm
  - One-step procedure
    - Fit the intensity  $\lambda$  and thinnings  $\{z^i(u)\}_i$  simultaneously from index, tranche and single name spreads
-

# Computational methods

- Semi-analytical characteristic function methods
- Monte Carlo simulation methods
  - Re-scaling algorithm
  - Exact algorithm
  - Approximate thinning algorithm
- Analytical time change methods

# Calculating the characteristic function

## Compensator

- Consider a non-explosive default counting process  $N$  with totally inaccessible event times  $T^k$
- By the Doob-Meyer theorem, there is a unique increasing continuous process  $A$  such that  $M = N - A$  is a local martingale; this process is called the **compensator** to  $A$
- Specifying the compensator is tantamount to specifying the evolution of  $N$ ; above we assumed that  $A$  has intensity  $\lambda$ :

$$A_t = \int_0^t \lambda_s ds$$

# Calculating the characteristic function

## Characteristic martingale

- Let  $\Psi(v) = 1 - e^{iv}$  and define

$$Z_t = e^{ivN_t + \Psi(v)A_t} \quad v \in \mathbb{R}$$

- We can show that  $Z$  satisfies the exponential equation

$$Z_t = 1 - \Psi(v) \int_0^t Z_{s-} dM_s$$

- It follows that  $Z$  is a local martingale
  - If for some  $T^*$  and constant  $c \geq 2$  the expectation of  $e^{cA_{T^*}}$  is finite, then for any  $T \leq T^*$  the stopped process  $Z^T = (Z_{t \wedge T})$  is a uniformly integrable martingale
-

# Calculating the characteristic function

Laplace transform

- For some  $s \leq T^*$ , consider the measure defined by  $Z = Z(v)$ :

$$\frac{dP^v}{dP} \Big|_{\mathcal{F}_t} = Z_t^s$$

- Denoting expectation under  $P^v$  by  $E^v$ , we define the **Laplace transform** of the compensator  $A$  under  $P^v$  by

$$\mathcal{L}_{t,s}(u, v) = E_t^v [e^{-u(A_s - A_t)}], \quad u \in \mathbb{C}, \quad v \in \mathbb{R}$$

- If  $P^v$  is a pricing measure and  $uA$  is the cumulative interest rate process, then  $\mathcal{L}_{t,s}(u, v)$  is the price of a zero coupon bond
  - Girsanov's theorem implies that  $N$  has  $P^v$ -compensator  $e^{iv} A$
-

# Calculating the characteristic function

Characteristic function

- **Theorem:** The characteristic function of  $N$  is given by

$$E_t \left[ e^{iv(N_s - N_t)} \right] = \mathcal{L}_{t,s}(\Psi(v), v)$$

- The calculation of the characteristic function reduces to the calculation of the Laplace transform of  $A$
  - Can be generalized to include stochastic interest rates that may depend on  $N$  (flight-to-quality)
-

# Calculating the characteristic function

Affine point process

- Consider a default process  $N$  with compensator

$$A_t = \int_0^t \lambda_s ds \quad \text{and} \quad \lambda = \Lambda_0 + \Lambda_1 X$$

for constants  $\Lambda_0$  and  $\Lambda_1$  such that  $\Lambda(X_t)$  is strictly positive

- The risk factor  $X$  follows the affine jump diffusion

$$dX_t = \mu(X_t)dt + \sigma(X_t)dW_t + \delta dL_t, \quad X_0 \in \mathbb{R}$$

where  $W$  is an independent Brownian motion and

- $\mu(x) = K_0 + K_1 x$ , where  $K_0$  and  $K_1$  are constants
- $\sigma^2(x) = H_0 + H_1 x$ , where  $H_0$  and  $H_1$  are constants

- Obvious generalization to multiple dimensions, time-dependent coefficients, additional jump terms
-

# Calculating the characteristic function

Affine point process

- We know that under  $P^v$ , the compensator of  $N$  is  $e^{iv}A$ , so we can use standard results on AJDs to get

$$\mathcal{L}_{t,s}(u, v) = E_t^v \left[ e^{-u \int_t^s \lambda_r dr} \right] = e^{\alpha(t) + \beta(t) X_t}$$

where  $\beta(t) = \beta(u, v, t, s)$  and  $\alpha(t) = \alpha(u, v, t, s)$  satisfy

$$\partial_t \beta(t) = u \Lambda_1 - K_1 \beta(t) - \frac{1}{2} H_1 \beta(t)^2 - e^{iv} \Lambda_1 (\theta(\delta \beta(t)) - 1)$$

$$\partial_t \alpha(t) = u \Lambda_0 - K_0 \beta(t) - \frac{1}{2} H_0 \beta(t)^2 - e^{iv} \Lambda_0 (\theta(\delta \beta(t)) - 1)$$

with boundary conditions  $\beta(s) = \alpha(s) = 0$ , and where, for  $\nu$  the distribution of the loss at default  $\ell^k$ ,

$$\theta(v) = \int e^{uz} d\nu(z)$$

# Calculating the characteristic function

Example: basic affine loss process

- Consider the special case  $H_0 = \Lambda_0 = 0$ ,  $H_1 = \sigma^2$ ,  $\Lambda_1 = 1$ ,  $K_0 = \kappa c$  and  $K_1 = -\kappa$  so that

$$d\lambda_t = \kappa(c - \lambda_t) dt + \sigma \sqrt{\lambda_t} dW_t + \delta dL_t$$

- The ODEs become

$$\partial_t \beta(t) = u + \kappa \beta(t) - \frac{1}{2} \sigma^2 \beta(t)^2 - e^{iv} (\theta(\delta \beta(t)) - 1)$$

$$\partial_t \alpha(t) = -c \kappa \beta(t)$$

- The characteristic function of  $N$  is

$$E_t[e^{iv(N_s - N_t)}] = \mathcal{L}_{t,s}(\Psi(v), v)$$

# Calculating the characteristic function

Example: basic affine loss process

- We can define new coefficient functions  $a$  and  $b$  such that

$$E_t[e^{iv(N_s - N_t)}] = \exp(a(v, t, s) + b(v, t, s)\lambda_t)$$

- The coefficient functions  $a(t) = a(v, t, s)$  and  $b(t) = b(v, t, s)$  solve the ordinary differential equations

$$\partial_t b(t) = 1 + \kappa b(t) - \frac{1}{2} \sigma^2 b(t)^2 - \theta^N(\delta b(t), v)$$

$$\partial_t a(t) = -c\kappa b(t)$$

with boundary conditions  $b(s) = a(s) = 0$ , and where, for  $\nu$  the distribution of the loss at default  $\ell^k$ ,

$$\theta^N(u, v) = e^{iv} \int e^{uz} d\nu(z)$$

# Calculating the characteristic function

Example: basic affine loss process

- Similarly, the characteristic function of the loss is

$$E_t[e^{iv(L_s - L_t)}] = \exp(a(v, t, s) + b(v, t, s)\lambda_t)$$

- The coefficient functions  $a(t) = a(v, t, s)$  and  $b(t) = b(v, t, s)$  solve the ordinary differential equations

$$\partial_t b(t) = 1 + \kappa b(t) - \frac{1}{2} \sigma^2 b(t)^2 - \theta^L(\delta b(t), v)$$

$$\partial_t a(t) = -c\kappa b(t)$$

with boundary conditions  $b(s) = a(s) = 0$ , and where, for  $\nu$  the distribution of the loss at default  $\ell^k$ ,

$$\theta^L(u, v) = \int e^{(u+iv)z} d\nu(z)$$

## Top down approach: key points

- The portfolio loss process is specified directly in terms of an intensity and a distribution for the loss at an event
    - Capture firm sensitivity to common factors and feedback
    - Capture dependence between default and recovery rates
    - Capture dependence between default and interest rates
  - The complete loss surface is described by one set of parameters, and so is the term structure of index and tranche spreads for all attachment points
  - Analytical, transform and simulation methods are available to efficiently calculate credit derivative prices
  - Single name hedges are generated by thinning the loss process
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## Notes

- The intensity based top down approach is based on the following papers:
    - Davis & Lo (2001): first application of idea
    - Errais, Giesecke & Goldberg (2006): specification of and valuation for affine loss processes
    - Giesecke & Goldberg (2005): analysis of random thinning
    - Giesecke (2007): characteristic function formula
    - Giesecke & Kim (2007): Monte Carlo simulation methods
    - Ding, Giesecke & Tomecek (2006): specification of and valuation for time-changed birth processes
    - Giesecke & Tomecek (2005): time changes in credit
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## Notes

- The **intensity based** top down approach is gaining popularity among academics and practitioners; recent references include
  - Arnsdorf & Halperin (2007)
  - Brigo, Pallavicini & Torresetti (2006)
  - Longstaff & Rajan (2006)
  - Lopatin & Misirpashaev (2007)
  - Tavella & Krekel (2006)

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